Statistics, Probability, Distributions, & Error Propagation

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Sample & Parent Populations

- Make measurements
 - $-x_{1}$
 - $-x_{2}$
 - In general do not expect $x_1 = x_2$
 - But as you take more and more measurements a pattern emerges in this sample
- With an infinite sample x_i , $i \in \{1...\infty\}$ we can
 - Expect a pattern to emerge with a characteristic value
 - Exactly specify the **distribution** of x_i
 - The hypothetical pool of all possible measurements is the *parent population*
 - Any finite sequence is the sample population

Histograms & Distributions

Simulated data 20 Histogram Sample x = 20.06Parent $\mu = 20.00$ represents the Sample s = 0.52 Parent $\sigma = 0.50$ **Jumber of measurements** occurrence or 15 frequency of discrete 10 measurements – Parent population 5 (dotted) Inferred parent distribution (solid) 18 21 22 19 20 Length (cm)

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Notation

- Parent distribution: Greek, e.g., μ
- Sample distribution: Latin, \bar{x}
 - To determine properties of the parent distribution assume that the properties of the sample distribution tend to those of the parent as *N* tends to infinity

Summation

• If we make *N* measurements, x_1 , x_2 , x_3 , etc. the sum of these measurements is

$$\sum_{i=1}^{N} x_i = x_1 + x_2 + x_3 + \dots + x_N$$

• Typically, we use the shorthand

$$\sum_{i=1}^{N} x_i = \sum x_i$$

Mean

• The mean of an experimental distribution is

$$\overline{x} = \frac{1}{N} \sum x_i$$

 The mean of the parent population is defined as

$$\mu = \lim_{N \to \infty} \left(\frac{1}{N} \sum x_i \right)$$

Median

• The median of the parent population $\mu_{1/2}$ is the value for which half of $x_i < \mu_{1/2}$

$$P(x_i < \mu_{1/2}) = P(x_i \ge \mu_{1/2}) = 1/2$$

• The median cuts the area under the probability distribution in half

Mode

- The mode is the most probable value drawn from the parent distribution
 - The mode is the most likely value to occur in an experiment
 - For a symmetrical distribution the mean, median and mode are all the same

Deviation

The deviation, d_i, of a measurement, x_i, from the mean is defined as

$$d_i = x_i - \mu$$

If μ is the true mean value the deviation is the error in x_i

Mean Deviation

• The mean deviation vanishes!

- Evident from the definition

$$\lim_{N \to \infty} \overline{d} = \lim_{N \to \infty} \left[\frac{1}{N} \sum (x_i - \mu) \right] = \lim_{\substack{N \to \infty \\ \mu}} \left[\frac{1}{N} \sum x_i \right] - \mu$$

Mean Square Deviation

 The mean square deviation is easy to use analytically and justified theoretically

$$\sigma^{2} = \lim_{N \to \infty} \left[\frac{1}{N} \sum (x_{i} - \mu)^{2} \right] = \lim_{N \to \infty} \left[\frac{1}{N} \sum x_{i}^{2} \right] - \mu^{2}$$

- σ^2 is also known as the *variance*
 - Derive this expression
 - Computation of σ^2 assumes we know μ

Population Mean Square Deviation

• The *estimate* of the standard deviation, *s*, from a sample population is

$$s^2 = \frac{1}{N-1} \sum \left(x_i - \overline{x} \right)^2$$

• The factor (*N*-1) is used instead of *N* to account for the fact that the mean must be derived from the data

Significance

- The mean of the sample is the best estimate of the mean of the parent distribution
 - The standard deviation, s, is characteristic of the uncertainties associated with attempts to measure μ

– But what is the uncertainty in μ ?

• To answer these questions we need probability distributions...

μ and σ of Distributions

- Define μ and σ in terms of the parent probability distribution P(x)
 - Definition of P(x)
 - Limit as $N \to \infty$
 - The number of observations *dN* that yield values between *x* and *x* + *dx* is

dN/N = P(x) dx

Expectation Values

 The mean, µ, is the expectation value of some quantity x

<x>

• The variance, σ^2 , is the expectation value of the deviation squared

Expectation Values

• For a discrete distribution, *N*, observations and *n* distinct outcomes

$$\mu = \lim_{N \to \infty} \frac{1}{N} \sum_{i=1}^{N} x_i$$

= $\lim_{N \to \infty} \frac{1}{N} \sum_{j=1}^{n} x_j n_{x_j}$ each x_j is a unique value
= $\lim_{N \to \infty} \frac{1}{N} \sum_{j=1}^{n} x_j NP(x_j)$
= $\lim_{N \to \infty} \sum_{j=1}^{n} x_j P(x_j)$

Expectation Values

• For a discrete distribution, *N*, observations and *n* distinct outcomes

$$\sigma^{2} = \lim_{N \to \infty} \frac{1}{N} \sum_{i=1}^{N} (x_{i} - \mu)^{2}$$
$$= \lim_{N \to \infty} \frac{1}{N} \sum_{j=1}^{n} (x_{j} - \mu)^{2} NP(x_{j})$$
$$= \lim_{N \to \infty} \sum_{j=1}^{n} \left[(x_{j} - \mu)^{2} P(x_{j}) \right]$$

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Expectation values

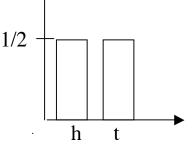
• The expectation value of any continuous function of *x*

$$\langle f(x) \rangle = \int_{-\infty}^{\infty} f(x) P(x) dx$$
$$\mu = \int_{-\infty}^{\infty} x P(x) dx$$
$$\sigma^{2} = \int_{-\infty}^{\infty} (x - \mu)^{2} P(x) dx$$

where
$$\int_{-\infty}^{\infty} P(x) dx = 1$$

Binomial Distribution

- Suppose we have two possible outcomes with probability *p* and *q* = 1-*p* – e.g., a coin toss, *p* = 1/2, *q* = 1/2
- If we flip *n* coins what is the probability of getting *x* heads?



- Answer is given by the Binomial Distribution

$$P(x;n,p) = C(n,x)p^{x}q^{n-x}$$

- C(n, x) is the number of combinations of n items taken x at a time = n!/[x!(n-x)!]

Binomial Distribution

• The expectation value

$$\mu = \sum_{x=0}^{n} x P(x;n,p)$$

= $\sum_{x=0}^{n} x C(n,x) p^{x} q^{n-x}$
= $\sum_{x=0}^{n} \left[x \frac{n!}{x!(n-x)!} p^{x} (1-p)^{n-x} \right] = np$

Poisson Distribution

- The Poisson distribution is the limit of the Binomial distribution when µ << n because p is small
 - The binomial distribution describes the probability *P*(*x*; *n*, *p*) of observing *x* events per unit time out of *n* possible events
 - Usually we don't know n or p but we do know μ

Poisson Distribution

• Suppose *p* << 1 then *x* << *n*

$$P(x;n,p) = \frac{n!}{x!(n-x)!} p^{x} (1-p)^{n-x}$$

$$\frac{n!}{(n-x)!} = n(n-1)(n-2)...(n-x-2)(n-x-1)$$

$$\approx n^{x} \text{ when } n \gg x$$

$$\frac{n!}{(n-x)!} p^{x} \approx (np)^{x} = \mu^{x}$$

$$(1-p)^{n-x} = (1-p)^{-x}(1-p)^{n} \approx 1 \times (1-p)^{n} \text{ since } p \ll 1$$

$$\lim_{p \to 0} (1-p)^{n} = \lim_{p \to 0} \left[(1-p)^{1/p} \right]^{\mu} = \left(e^{-1}\right)^{\mu} = e^{-\mu}$$

$$P(x,\mu) = \frac{\mu^{x}}{x!} e^{-\mu}$$

Poisson Distribution

• The expectation value of x is

$$\langle x \rangle = \sum_{x=0}^{\infty} x P(x,\mu) = \sum_{x=0}^{\infty} x \frac{\mu^x}{x!} e^{-\mu} = \mu$$

• Expectation value of $(x-\mu)^2$

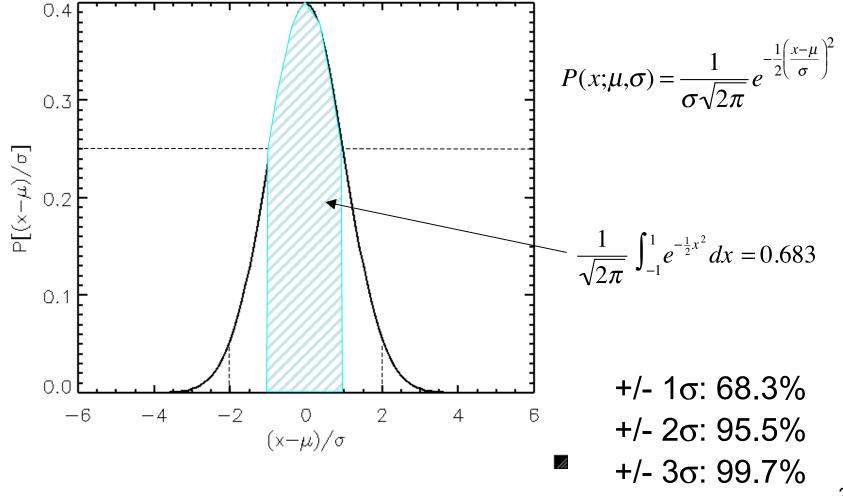
$$\sigma^{2} = \left\langle \left(x - \mu \right)^{2} \right\rangle = \sum_{x=0}^{\infty} \left(x - \mu \right)^{2} \frac{\mu^{x}}{x!} e^{-\mu} = \mu$$

Gaussian or Normal Distribution

 The Gaussian distribution is an approximation to the binomial distribution for large n and large np

$$P(x;\mu,\sigma) = \frac{1}{\sigma\sqrt{2\pi}}e^{-\frac{1}{2}\left(\frac{x-\mu}{\sigma}\right)^2}$$

Gaussian or Normal Distribution



- Suppose I have two sets of measurements, *a_i*, and *b_i*
 - A derived quantity $c_i = a_i + b_i$
 - What is the relation between the means and standard deviations of a_i and b_i and c_i
 - Suppose we have the same number of observations N of a_i and b_i

$$N = N_a = N_b$$

$$\overline{a} = \frac{1}{N} \sum a_i \quad \overline{b} = \frac{1}{N} \sum b_i$$

$$\overline{c} = \frac{1}{N} \sum c_i \quad s_c^2 = \frac{1}{N-1} \sum (c_i - \overline{c})^2$$

$$c_i = a_i + b_i$$

$$\overline{c} = \frac{1}{N} \sum (a_i + b_i) = \frac{1}{N} \sum a_i + \frac{1}{N} \sum b_i$$

$$= \overline{a} + \overline{b}$$

$$s_c^2 = \frac{1}{N-1} \sum (c_i - \overline{c})^2, \quad \overline{c} = \overline{a} + \overline{b}$$

$$s_c^2 = \frac{1}{N-1} \sum \left[a_i + b_i - (\overline{a} + \overline{b}) \right]^2$$

$$= \frac{1}{N-1} \sum \left[(a_i + b_i)^2 - 2(a_i + b_i)(\overline{a} + \overline{b}) + (\overline{a} + \overline{b})^2 \right]$$

$$= \frac{1}{N-1} \sum \left[a_i^2 + b_i^2 + 2a_i b_i - 2(a_i \overline{a} + a_i \overline{b} + b_i \overline{a} + b_i \overline{b}) + (\overline{a})^2 + 2\overline{a}\overline{b} + (\overline{b})^2 \right]$$

$$= \frac{N}{N-1}\overline{a^2} + \frac{N}{N-1}\overline{b^2} + \frac{2}{N-1}\sum a_i b_i - \frac{N}{N-1}(\overline{a})^2 - \frac{2N}{N-1}\overline{a}\overline{b} - \frac{N}{N-1}(\overline{b})^2$$

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$$s_{c}^{2} = \frac{1}{N-1} \sum (c_{i} - \overline{c})^{2}, \quad \overline{c} = \overline{a} + \overline{b}$$

$$= \frac{N}{N-1} \overline{a^{2}} + \frac{N}{N-1} \overline{b^{2}} + \frac{2}{N-1} \sum a_{i} b_{i} - \frac{N}{N-1} (\overline{a})^{2} - \frac{2N}{N-1} \overline{a} \overline{b} - \frac{N}{N-1} (\overline{b})^{2}$$

$$= \underbrace{\frac{N}{N-1} \left[\overline{a^{2}} - (\overline{a})^{2} \right]}_{s_{a}^{2}} + \underbrace{\frac{N}{N-1} \left[\overline{b^{2}} - (\overline{b})^{2} \right]}_{s_{b}^{2}} + \underbrace{\frac{2N}{N-1} \left(\overline{a} \overline{b} - \overline{a} \overline{b} \right)}_{2s_{ab}^{2}}$$

$$s_{c}^{2} = s_{a}^{2} + s_{b}^{2} + 2s_{ab}^{2}$$

- The term s_{ab}^2 is the covariance
 - Murphy's law factor
 - $-s_{ab}$ can be negative, zero or positive

Combining Two Uncorrelated Observations

• When a and b are uncorrelated the covariance is zero

$$s_{ab}^2 = \frac{1}{N-1} \sum \left(a_i - \overline{a} \right) \left(b_i - \overline{b} \right) = 0$$

$$s_c^2 = s_a^2 + s_b^2$$

- The variance of c is the sum of the variances of a and b
- This demonstrates the fundamentals of error propagation

 Suppose we want to determine x which is a function of measured quantities, u, v, etc.

$$x = f(u, v, ...)$$

Assume that

$$\overline{x} = f(\overline{u}, \overline{v}, \dots)$$

 The uncertainty in x can be found by considering the spread of the values of x resulting from individual measurements, u_i, v_i, etc.,

$$x_i = f(u_i, v_i, \dots)$$

• In the limit of $N \rightarrow \infty$ the variance of x

$$\sigma_x^2 = \lim_{N \to \infty} \frac{1}{N} \sum_i (x_i - \bar{x})^2$$

Taylor expand the deviation (N→∞ assumed

$$\begin{aligned} x_i - \overline{x} &= \left(u_i - \overline{u}\right) \frac{\partial f}{\partial u}\Big|_{\overline{u}} + \left(v_i - \overline{v}\right) \frac{\partial f}{\partial v}\Big|_{\overline{v}} + \dots \\ \sigma_x^2 &= \frac{1}{N} \sum_i \left[\left(u_i - \overline{u}\right) \frac{\partial f}{\partial u}\Big|_{\overline{u}} + \left(v_i - \overline{v}\right) \frac{\partial f}{\partial v}\Big|_{\overline{v}} + \dots \right]^2 \\ &= \frac{1}{N} \sum_i \left[\left(u_i - \overline{u}\right)^2 \left(\frac{\partial f}{\partial u}\right)_{\overline{u}}^2 + \left(v_i - \overline{v}\right)^2 \left(\frac{\partial f}{\partial v}\right)_{\overline{v}}^2 + 2\left(u_i - \overline{u}\right) \left(v_i - \overline{v}\right) \frac{\partial f}{\partial u}\Big|_{\overline{u}} \frac{\partial f}{\partial v}\Big|_{\overline{v}} \dots \right] \end{aligned}$$

$$\sigma_x^2 = \frac{1}{N} \sum_i \left[\left(u_i - \overline{u} \right)^2 \left(\frac{\partial f}{\partial u} \right)_{\overline{u}}^2 + \left(v_i - \overline{v} \right)^2 \left(\frac{\partial f}{\partial v} \right)_{\overline{v}}^2 + 2 \left(u_i - \overline{u} \right) \left(v_i - \overline{v} \right) \frac{\partial f}{\partial u} \Big|_{\overline{u}} \frac{\partial f}{\partial v} \Big|_{\overline{v}} \dots \right]$$

$$= \frac{1}{N} \sum_i \left(u_i - \overline{u} \right)^2 \left(\frac{\partial f}{\partial u} \right)_{\overline{u}}^2 + \frac{1}{N} \sum_i \left(v_i - \overline{v} \right)^2 \left(\frac{\partial f}{\partial v} \right)_{\overline{v}}^2 + \frac{2}{N} \sum_i \left(u_i - \overline{u} \right) \left(v_i - \overline{v} \right) \frac{\partial f}{\partial u} \Big|_{\overline{u}} \frac{\partial f}{\partial v} \Big|_{\overline{v}} + \dots$$

$$\sigma_x^2 = \sigma_u^2 \left(\frac{\partial f}{\partial u} \right)_{\overline{u}}^2 + \sigma_v^2 \left(\frac{\partial f}{\partial v} \right)_{\overline{v}}^2 + 2\sigma_{uv}^2 \left(\frac{\partial f}{\partial u} \right)_{\overline{u}}^2 + \dots$$

Examples of Error Propagation

• Suppose a = b + c

- We know that

$$\overline{a} = \overline{b} + \overline{c}$$
$$\sigma_a^2 = \sigma_b^2 + \sigma_c^2$$

assuming that the covariance is 0

• What about *a* = *b/c*?

Examples of Error Propagation

• Suppose *a* = *b/c*?

$$\overline{a} = \overline{b} / \overline{c}$$

and

$$\sigma_a^2 = \sigma_b^2 \left(\frac{\partial a}{\partial b}\right)_{\bar{b}}^2 + \sigma_c^2 \left(\frac{\partial a}{\partial c}\right)_{\bar{c}}^2 + 2\sigma_{bc}^2 \frac{\partial a}{\partial b}\Big|_{\bar{b}} \frac{\partial a}{\partial c}\Big|_{\bar{c}} + \dots$$
$$\sigma_a^2 = \sigma_b^2 \frac{1}{c^2} + \sigma_c^2 \left(\frac{b}{c^2}\right)^2$$
or
$$\frac{\sigma_a^2}{a^2} = \frac{\sigma_b^2}{b^2} + \frac{\sigma_c^2}{c^2}$$

assuming that the covariance is 0

Error of the Mean

 Suppose we have N measurements, x_i with uncertainties characterized by s_i

$$\overline{x} = \frac{1}{N} \left(x_1 + x_2 + x_3 + \dots + x_N \right) = \frac{1}{N} \sum_i x_i$$

$$s_x^2 = s_1^2 \left(\frac{\partial \overline{x}}{\partial x_1} \right)_{\overline{x}}^2 + s_2^2 \left(\frac{\partial \overline{x}}{\partial x_2} \right)_{\overline{x}}^2 + s_3^2 \left(\frac{\partial \overline{x}}{\partial x_3} \right)_{\overline{x}}^2 + \dots + s_N^2 \left(\frac{\partial \overline{x}}{\partial x_N} \right)_{\overline{x}}^2$$

$$= \sum_i s_i^2 \left(\frac{\partial \overline{x}}{\partial x_i} \right)_{\overline{x}}^2$$

assuming that the covariance is 0

Error of the Mean

 Suppose the errors on all points are equal so that s_i = s

$$s_{\overline{x}}^{2} = \sum_{i} s_{i}^{2} \left(\frac{\partial \overline{x}}{\partial x_{i}}\right)_{\overline{x}}^{2}$$
$$\frac{\partial \overline{x}}{\partial x_{i}} = \frac{\partial}{\partial x_{i}} \left(\frac{1}{N} \sum_{j} x_{j}\right) = \frac{1}{N} \qquad \qquad \frac{\partial x_{j}}{\partial x_{i}} = \delta_{ij}$$
$$s_{\overline{x}}^{2} = \sum_{i} s^{2} \left(\frac{1}{N}\right)^{2}$$
$$= \frac{s^{2}}{N}$$

Examples of Error Propagation

- What happens when $m = -2.5 \log_{10}(F/F_0)$?
 - What is the error in *m*?

$$m = -2.5 \log_{10} \left(F/F_0 \right)$$

and

$$\sigma_m^2 = \sigma_F^2 \left(\frac{\partial m}{\partial F}\right)_{\overline{F}}^2$$
$$\sigma_m^2 = \sigma_F^2 \left(\frac{2.5}{F \log(10)}\right)^2$$
$$\sigma_m^2 = (1.087)^2 \left(\frac{\sigma_F}{F}\right)^2$$